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Application of Sequential Design of Experiments (SDoE) to Large Pilot-Scale Solvent-Based CO₂ Capture Process at Technology Centre Mongstad (TCM)

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Abstract

The United States Department of Energy's Carbon Capture Simulation for Industry Impact (CCSI²) program has developed a framework for sequential design of experiments (SDoE) that aims to maximize knowledge gained from budget- and schedule-limited pilot scale testing. SDoE was applied to the planning and execution of campaigns for testing CO₂ capture systems at pilot-scale in order to optimally allocate resources available for the testing. In this methodology, a stochastic process model is developed by quantifying the parametric uncertainty in submodels of interest; for a solvent-based CO₂ capture system, these may include physical properties and equipment performance submodels (e.g., mass transfer, interfacial area). This uncertainty is propagated through the full process model, over variable operating conditions, for estimating the resulting uncertainty in key model outputs (e.g., percentage of CO₂ capture, solvent regeneration energy requirement). In developing a data collection plan, the predicted output uncertainty is incorporated into an algorithm that seeks simultaneously to select process operating conditions for which the predicted uncertainty is relatively high and to ensure that the entire space of operation is well represented. This test plan is then used to guide operation of the pilot plant at varying steady-state conditions, with resulting process data incorporated into the existing model using Bayesian inference to refine parameter distributions. The updated stochastic model, with reduced parametric uncertainty from data collected, is then used to guide additional data collection, thus the sequential nature of the experimental design.

The SDoE process was implemented at the pilot test unit (12 MWe in scale) at Norway's Technology Centre Mongstad (TCM) in a summer 2018 test campaign with aqueous monoethanolamine (MEA). During the test campaign, the varied operating conditions included the flowrates of circulated solvent, flue gas, and reboiler steam and the CO₂ concentration in the flue gas. The process data were used to update probability distributions of mass transfer and interfacial area parameters of a stochastic process model developed by the CCSI² team. Two iterations of the SDoE process were executed, resulting in the uncertainty in model predicted CO₂ capture percentage decreasing by an average of $58.0 \pm 4.7\%$ over the full input space of interest. This work demonstrates the

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potential of the SDoE process for model refinement through reduction in process model parametric uncertainty, and ultimately risk in scale-up, in CO₂ capture technology performance.

Keywords: post-combustion carbon capture; pilot-scale testing; uncertainty quantification; design of experiments

1. Introduction

The United States Department of Energy's Carbon Capture Simulation for Industry Impact (CCSI²) program is a collaboration of national laboratories, universities, and industrial organizations that provides research and development support for novel CO₂ capture technologies with the objective of reducing risk and accelerating their commercialization. These efforts involve continuing advancements in and applications of the open-source toolset¹ developed as part of its predecessor project, the Carbon Capture Simulation Initiative (CCSI). The CCSI Toolset includes a suite of computational tools and models with the overarching goal of accelerating the development, deployment, and scale-up of CO₂ capture technologies. The toolset includes a rigorous process model, implemented in Aspen Plus[®], of the aqueous monoethanolamine (MEA) solvent system, which is the industrial standard for solvent-based CO₂ capture. This model includes quantification of parametric uncertainty for solvent physical property models such as viscosity, density, and surface tension [1], the thermodynamic framework [2], and packing-specific models such as mass transfer, interfacial area, and hydraulics [3]. These submodels combine with a full process model that was validated with process data from the 0.5 MWe pilot test unit at the National Carbon Capture Center (NCCC) in 2014 [4]. In 2017, an additional test campaign for the aqueous MEA system was held at NCCC, incorporating the CCSI² framework for SDoE. In this methodology, the existing process model is leveraged to inform collection of data that are subsequently used to refine the model and modify the test plan accordingly [5,6]. Over two iterations of the SDoE process, parametric distributions for process submodels were refined through experimental observations of absorber CO₂ capture percentage, resulting in an average uncertainty reduction of approximately 50% for the model prediction of CO₂ capture percentage throughout the input space of interest.

The CCSI aqueous MEA process model was scaled up to 12 MWe for consistency with the pilot test unit at Norway's Technology Centre Mongstad (TCM) and was used in the planning and execution of a test campaign at TCM in summer 2018. TCM is one of the world's largest facilities for testing carbon capture technologies, and previous test results with the MEA solvent system have been reported in the open literature [7-12], including variation in many process variables and both steady-state and dynamic operation. The pilot plant at TCM notably has two sources of flue gas: combined cycle gas turbine (CCGT) based heat and power plant (CHP), with ~3.5 vol% CO₂, and residual fluidized catalytic cracker (RFCC) unit, with ~13-14 vol% CO₂. The TCM plant also contains two stripper columns, each designed for process operation with one of the flue gas sources. This work focused on collecting additional data for the MEA process at TCM with variation in the flowrates of solvent, flue gas, and reboiler steam, the concentration of CO₂ in the flue gas, the packing height of the absorber, and the stripper configuration. During the first three weeks of the test campaign, which are the primary focus of this paper, the SDoE framework was used to guide the collection of process data using the existing MEA process model and multiple test objectives. The data were used to update the model by refining the distributions of parameters in the mass transfer and interfacial area submodels, ultimately resulting in a reduction of predicted uncertainty in the CO₂ capture percentage from $10.5 \pm 1.5\%$ to $4.4 \pm 0.4\%$, or an average reduction of $58.0 \pm 4.7\%$, over the full input space of interest. In the final two weeks of the campaign, data were collected for a modified process configuration in which the packing height of the absorber was reduced to 18 meters, and eventually 12 meters, and the stripper configuration was modified so that a fraction (~20%) of the rich solvent exiting the absorber bypasses the lean-rich heat exchanger and is heated in the water wash of the stripper. This work, along with the previous test campaign at NCCC, demonstrates the potential of the SDoE methodology for refining predictions of stochastic process models through strategic data collection. The reduction of model uncertainty effectively reduces expected risk in process design and operation, thus improving confidence when predicting process performance and conducting economic analyses.

¹ Available at <https://github.com/CCSI-Toolset/>

2. Methodology

2.1. SDoE Methodology

The SDoE process developed by CCSI² uses a stochastic model, with parametric uncertainty quantified in the submodels, to inform collection of process data in order to maximize the value of data obtained during a test campaign. Moreover, it provides a framework for directly reducing uncertainty in model prediction of capture rates based on new process knowledge gained from data collection. The SDoE process is represented schematically in Fig. 1.

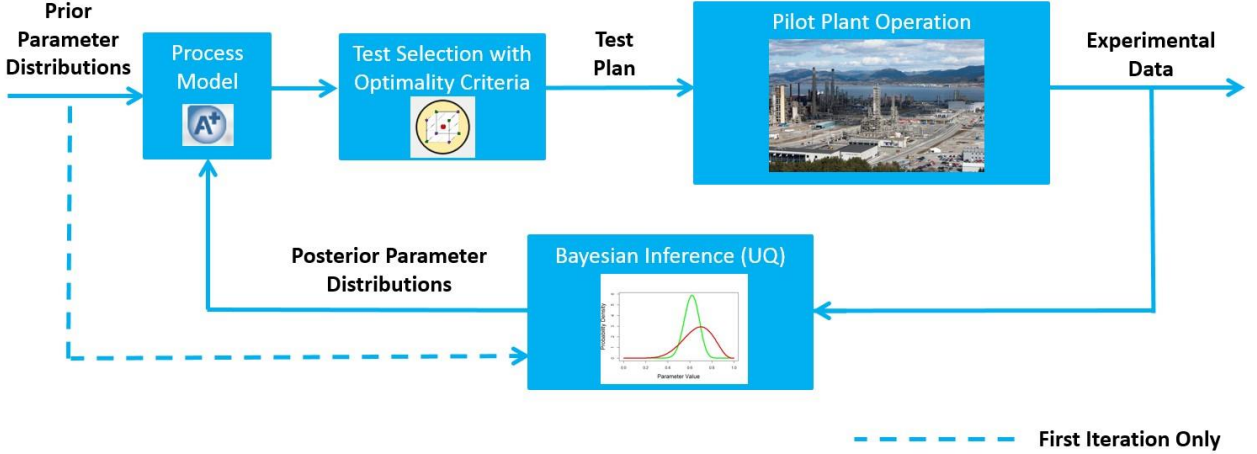


Fig. 1. Schematic representation of SDoE methodology

As shown in Figure 1, *a priori* probability distributions of submodel parameters are propagated through a process model, denoted as $y(\mathbf{x}, \boldsymbol{\theta}, \boldsymbol{\theta}^*)$, where \mathbf{x} is a set of input variables that can be manipulated in plant operation and $\boldsymbol{\theta}$ and $\boldsymbol{\theta}^*$ are sets of model parameters. These sets differ in that $\boldsymbol{\theta}$ contains parameters for which uncertainty is reduced over the course of executing the SDoE methodology whereas $\boldsymbol{\theta}^*$ has parameters with irreducible uncertainty based on the type of data collected. For the example of a solvent-based CO₂ capture system, $\boldsymbol{\theta}$ includes process specific parameters from submodels such as mass transfer or interfacial area that may be informed through collection of plant data (e.g. CO₂ capture percentage in the absorber column). Conversely, $\boldsymbol{\theta}^*$ includes parameters from physical property submodels, for which uncertainty is readily estimated through direct measurements of the corresponding properties and cannot reasonably be informed from plant level data. If the process model y is sufficiently complex, it may be necessary to replace it with a surrogate model, denoted as $\hat{y}(\mathbf{x}, \boldsymbol{\theta}, \boldsymbol{\theta}^*)$, developed and validated over the full input space. For a given point in the input space, a confidence interval for the model prediction are computed by propagating the uncertainty in the full set of parameters ($\boldsymbol{\theta}^T = [\boldsymbol{\theta} \quad \boldsymbol{\theta}^*]$) through the surrogate model. The 95% confidence interval, estimated by taking a sample of size M over the full parameter space ($\boldsymbol{\theta}^{T,(j)}, \forall j = 1, \dots, M$), is given as:

$$CI|_{\mathbf{x}^{(i)}} = F_{0.975}(\{\hat{y}(\mathbf{x}^{(i)}, \boldsymbol{\theta}^{T,(1)}), \dots, \hat{y}(\mathbf{x}^{(i)}, \boldsymbol{\theta}^{T,(M)})\}) - F_{0.025}(\{\hat{y}(\mathbf{x}^{(i)}, \boldsymbol{\theta}^{T,(1)}), \dots, \hat{y}(\mathbf{x}^{(i)}, \boldsymbol{\theta}^{T,(M)})\}) \quad (1)$$

where $\{\hat{y}(\mathbf{x}^{(i)}, \boldsymbol{\theta}^{T,(1)}), \dots, \hat{y}(\mathbf{x}^{(i)}, \boldsymbol{\theta}^{T,(M)})\}$ is the set of values of an output variable calculated from propagating all of the individual $\boldsymbol{\theta}^{T,(j)}$ through the surrogate model and F_k represents the k^{th} percentile of this set. The values of $CI|_{\mathbf{x}^{(i)}}$ for individual $\mathbf{x}^{(i)}$ are considered in the test selection method; the specific optimality criterion used in this work is G-optimality [13], which minimizes the maximum prediction variance. This aim targets experimental settings $\mathbf{x}^{(i)}$ for which the predicted uncertainty (i.e., $CI|_{\mathbf{x}^{(i)}}$) is relatively large, so that the collection of data at these settings represents high potential for uncertainty reduction. Moreover, the algorithm used in this work for test selection simultaneously seeks to ensure that the full input space is well-represented in the test plan, balancing good representation of design points throughout the region while making locations with large confidence interval widths more likely to be selected.

The test plan is then implemented by running the plant accordingly, resulting in collection of experimental data (denoted Z). The data are incorporated into a Bayesian inference framework, using the PSUADE² software package. For model parameters of fixed uncertainty, a sample $(\theta^{*(j)}; \forall j = 1, \dots, N)$ is drawn from their probability distribution $P(\theta^*)$. For each sample point $\theta^{*(j)}$, a posterior distribution for the remaining parameters (θ) is calculated:

$$\pi_j(\theta|Z, \theta^{*(j)}) \propto P(\theta)L(Z|\theta, \theta^{*(j)}) \quad (2)$$

and given in the form of a set of sample points. Here, $L(Z|\theta, \theta^{*(j)})$ represents the likelihood (some metric used to express the distance between simulation predictions and experimental data) of observing a set of experimental data (Z) conditioned on the values of the parameters, $P(\theta)$ the prior distribution of the parameters for which uncertainty is updated, and $\pi_j(\theta|Z, \theta^{*(j)})$ the posterior distribution of θ conditioned on the observed experimental data and the value of θ^* for sample j . The overall posterior distribution $\pi(\theta|Z, \theta^{*(j)})$ is obtained through the process of marginalization, combining all individual $\pi_j(\theta|Z, \theta^{*(j)})$. The updated stochastic model, with refined estimates of parameter uncertainties, is then used to re-calculate $CI|_{x^{(t)}}$ throughout the input space. For all subsequent iterations of SDoE, the prior distribution $P(\theta)$ is replaced by the posterior distribution $\pi_j(\theta|Z, \theta^{*(j)})$ from the previous iteration.

2.2. Overview of TCM Test Campaign

The TCM test campaign ran for five weeks in summer 2018, in five distinct test phases as outlined in Table 1.

Table 1. Phases of MEA test campaign at TCM

Phase No.	Absorber Packing Height (m)	CO ₂ in Flue Gas (vol%)	No. of Data Sets	Stripper Configuration	Description of SDoE Criterion
1	24	8	14	Simple	Space-Filling Design
2	24	8 & 10	10	Simple	Selection of points with optimal economic performance
3	24	8 & 10	41	Simple	Sequential SDoE targeting uncertainty reduction
4	18	10	14	With Bypass	Minimization of specific reboiler duty (SRD)
5	12	10	19	With Bypass	Minimization of SRD

In the first three phases of the campaign, the absorber column was operated with all three packing beds (total height of 24 meters). A conventional stripper configuration was used in which the full amount of rich solvent exiting the absorber is heated in the lean-rich heat exchanger and sent to the top of the stripper column. Throughout the test campaign, flue gas from the CCGT plant (3.5 vol% CO₂) was combined with recycle of the captured CO₂, increasing the flue gas concentration to 8 or 10 vol% CO₂ as required by the test plan. Due to the increased CO₂ concentration in the flue gas, and the corresponding increase in the required solvent circulation rate for capturing CO₂, the larger stripper intended for use with RFCC flue gas was used during this campaign in lieu of the smaller stripper intended for CCGT flue gas. In Phases 4-5, the packing height of the absorber was reduced by changing the number of beds and the stripper configuration was modified so that approximately 20% of the rich solvent exiting the absorber column bypassed the lean-rich heat exchanger and was instead heated with hot vapor leaving the top of the stripper. This portion of the test campaign, also guided with use of the process model, was focused on identifying the optimal solvent circulation for minimizing the specific reboiler duty for the process. Other process variables were fixed for this portion of the test campaign, including a flue gas flowrate of 50,000 sm³/hr with 10 vol% CO₂ and 85% CO₂ capture. For the purpose of brevity, the details of Phases 4-5 are not included in this paper.

² Problem Solving Environment for Uncertainty Analysis and Design Exploration (<https://computing.llnl.gov/projects/psuade-uncertainty-quantification>)

The first three phases differed in the choice of criteria used for developing the test plan. Phase 1 used a space-filling design to ensure that the process model predicts key outputs such as CO₂ capture and specific reboiler duty with reasonable accuracy. Phase 2 focused on collecting data in regions where the model predicts optimal economic performance. Phase 3 focused on collecting data to target regions where the model predicts high uncertainty based on the SDoE procedure in Section 2.1. For all three phases, a desired region of process operation was established based on ranges of operation for flue gas flowrate (G), CO₂ capture percentage (CAP), CO₂ loading in the lean solvent entering the absorber column (α_{lean}), and the volume fraction of CO₂ in the flue gas (y_{CO_2}), with ranges:

$$G \in [36,000 - 75,000] \text{ kg/hr} \quad (3a)$$

$$CAP \in [80 - 95] \% \quad (3b)$$

$$\alpha_{lean} \in [0.10 - 0.25] \text{ mol CO}_2/\text{mol MEA} \quad (3c)$$

$$y_{CO_2} \in \{0.08, 0.10\} \quad (3d)$$

The first three variables are treated as continuous whereas the CO₂ fraction in the flue gas is treated as a categorical variable with two process operation levels. For each value of CO₂ fraction, a test set consisting of candidate experiments with a unique combination of variables $\{G, CAP, \alpha_{lean}\}$ is generated by sampling independently from uniform distributions for each variable with upper and lower limits based on the ranges given in Eq. 3. An Aspen Plus simulation is run for each point in the candidate set to estimate the corresponding values of lean (L_{lean}) and rich (L_{rich}) solvent flowrate, steam flowrate (S) and mass of CO₂ captured (\dot{m}_{CAP}). To be included in the final candidate set, a point must satisfy the following conditions based on operational limits for the TCM plant:

$$\dot{m}_{CAP} < 8,000 \text{ kg/hr} \quad (4a)$$

$$S < 14,000 \text{ kg/hr} \quad (4b)$$

Separate candidate sets (for $y_{CO_2} = 0.08$ and $y_{CO_2} = 0.10$) were developed using a space-filling approach based on the input vector $\mathbf{x} = [G \ S \ \alpha_{lean}]$. These candidate sets were used in Phases 1 and 3, although Phase 1 used a space-filling design on the model input space while Phase 3 incorporated the predicted uncertainty in the model output, using the methodology described in Section 2.1. Moreover, only the candidate set for 8 vol% CO₂ in flue gas was implemented during Phase 1 of the test campaign due to time considerations.

Phase 2, however, was designed based on an optimization problem of the form:

$$\min_{\mathbf{x}} f(\mathbf{x}) = \frac{CAPEX \left(\frac{A}{P}, i, n \right) + OPEX}{\dot{m}_{CAP}} \quad (5a)$$

$$\left(\frac{A}{P}, i, n \right) = \frac{i(1+i)^n}{(1+i)^n - 1} \quad (5b)$$

$$\mathbf{x} = \begin{bmatrix} L_{lean} \\ G \\ \alpha_{lean} \end{bmatrix} \quad (5c)$$

$$\text{subject to:} \\ \mathbf{x}^L \leq \mathbf{x} \leq \mathbf{x}^U \quad (5d)$$

$$h(\mathbf{x}) = 0 \quad (5e)$$

$$g(\mathbf{x}) \leq 0 \quad (5f)$$

The objective function is the ratio of the equivalent annual operating cost (EAOC) associated with the CO₂ capture to the mass of CO₂ captured. The EAOC is the sum of the capital cost (CAPEX) multiplied by an annuity factor $\left(\frac{A}{P}, i, n \right)$ and the operating cost (OPEX). Within the annuity factor, $\frac{A}{P}$ is the ratio of annuity to present value, i is the interest rate, and n is the number of years. The vector of decision variables is denoted as \mathbf{x} with lower and upper bounds \mathbf{x}^L and \mathbf{x}^U . The equality constraints denoted by $h(\mathbf{x})$ includes heat and material balances, and the inequality constraints

denoted by $g(x)$ includes the constraints for process operation listed in Eq. 4. The optimization was performed separately for the cases with 8 and 10 vol% CO₂ in flue gas. In addition the optimal points in the test plan, additional test points near the optimal points were included. The space surrounding the optimal point can be represented by a cube created by perturbing the input variable values by a chosen amount (10% for this study) from their estimated optimal values. A design that permutes each factor away from this estimated optimum one at a time would require seven test points, or six for the center of each face of the cube (if each factor is manipulated one at a time) and one for the center (optimal) point. As shown in Fig. 2., the design size was reduced to five by considering a fractional factorial structure, which also allows exploration of potential interactions between input factors around the optimum [14].

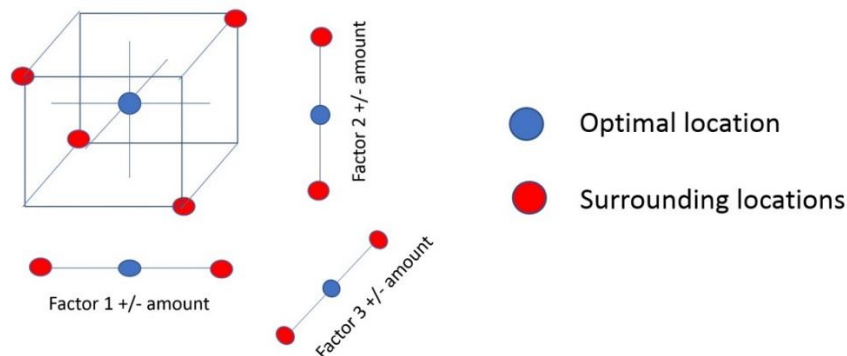


Fig. 2. Space-filling in region around optimal point for Phase 2 test plan

Since two levels of y_{CO_2} were included in the analysis, the reduction of the overall number of points required for the Phase 2 test plan from 14 to 10 was highly beneficial due to the limited amount of time available for the test campaign.

3. Results

3.1 Phase 1

In Phase 1, the test plan was developed using a minimax space-filling methodology [15] to provide an initial data set that was well-representative of the process. For all testing in this phase, the CO₂ concentration in the flue gas was fixed at 8 vol%. The set of input variables included in the test matrix differs from that used for space-filling design in that the input variables for the space-filling design were chosen for modeling convenience whereas the input variables in the test matrix were those directly manipulated in the plant operation. In developing the test matrix, the Aspen simulation was used to estimate the rich solvent flowrate and the flue gas flowrate was converted from mass to volumetric units. The test matrix, which was organized in terms of increasing flue gas flowrate for ease of process operation, for Phase 1 is given in Table 2.

Table 2. Test matrix for Phase 1 design of MEA test campaign at TCM

Test	Rich Solvent Flowrate (kg/hr)	Flue Gas Flowrate (Sm ³ /hr)	Steam Flowrate (kg/hr)	CO ₂ Capture Percent Estimate
1A	55,300	31,800	5,500	86.1
1B	54,200	36,000	7,200	88.0
1C	92,100	37,300	7,400	92.5
1D	81,400	43,800	7,700	84.9
1E	81,300	45,900	8,900	93.4
1F	120,800	53,700	10,700	92.2
1G	88,900	56,500	12,100	90.4
1H	90,300	57,100	9,800	82.7

When obtaining data for test cases 1A-1B, it was noted that the CO₂ capture percentage was substantially lower than the model predictions. This discrepancy was attributed to solvent maldistribution, or uneven flow through the packing, in the RFCC stripper column, resulting in inefficient performance of the column. This stripper was designed to operate at a solvent flowrate of approximately 200,000 kg/hr, or almost four times higher than the solvent flowrate in cases 1A-1B. Therefore, the lean solvent loading for these test runs was substantially higher than

that predicted by the model, and the CO₂ capture percentage lower. This issue was addressed by dividing each subsequent test run into two intervals with distinct operating goals, so that two data sets were collected for test runs 1C-1H. First, the test was executed with the value of steam flowrate specified in the original test matrix. Upon achieving the steady-state, the steam flowrate was manipulated to match the estimated value of CO₂ capture. Parity plots for the model prediction of CO₂ capture percentage in the absorber and steam requirement in the stripper are given in Figure 3.

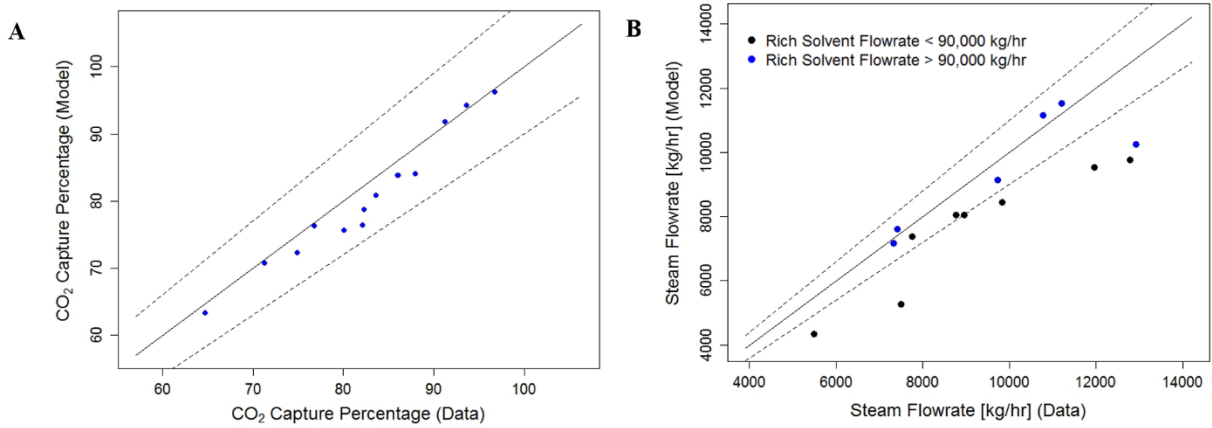


Figure 3. Parity plots for (A) CO₂ capture percentage and (B) steam flowrate required for test runs performed in Phase 1. Dashed lines represent $\pm 10\%$ error.

The original deterministic model, or the model in which all parameters are fixed as point values, predicts the CO₂ capture with a range of $\pm 10\%$ error for all test runs. The average error for CO₂ capture percentage is $-2.51 \pm 2.29\%$, with the negative error indicating that the model generally underpredicts the data. The model predicts stripper steam requirement with an average error of $-10.83 \pm 10.82\%$, although the error is notably higher for cases in which the solvent flowrate is below 90,000 kg/hr (average error of $-16.43 \pm 8.49\%$) than when it is higher than 90,000 kg/hr (average error of $-3.67 \pm 9.29\%$). This discrepancy is likely due to liquid maldistribution in the stripper column, as discussed previously. The results obtained in the first phase of the test campaign demonstrated that the initial process model was sufficiently accurate to proceed with the sequential experimental design in subsequent stages.

3.3 Phases 2-3

During the test campaign, data for Phases 2-3 were collected simultaneously and used to update the model parameter distributions through Bayesian inference. The majority of the data for Phase 2 were actually collected after those for Phase 3 due to scheduling convenience. The optimization problem described in Eq. 5 was implemented separately for the 8 and 10 vol% CO₂ cases, and used to develop the test matrix given in Table 3.

Table 3. Test matrix for Phase 2 design of MEA test campaign at TCM

Test	Rich Solvent Flowrate (kg/hr)	Flue Gas Flowrate (Sm ³ /hr)	Steam Flowrate (kg/hr)	CO ₂ in Flue Gas (vol%)
2A	107,800	40,800	10,700	10
2B	107,100	44,100	10,300	8
2C	107,100	44,100	12,500	8
2D	97,400	49,000	11,400	8
2E	87,700	53,900	10,300	8
2F	87,700	53,900	12,500	8
2G	97,000	44,900	11,800	10
2H	97,000	44,900	9,600	10
2I	118,600	36,700	9,600	10

In Table 3, the optimal points determined from solving separate optimization problems (Eq. 5) for the $y_{CO_2} = 0.08$ (2A) and $y_{CO_2} = 0.10$ (2D) cases are highlighted, and additional test points were selected by perturbing the variables

by $\pm 10\%$ from the optimal values. Parity plots for the model prediction of CO₂ capture percentage in the absorber and steam requirement in the stripper are given in Fig. 4.

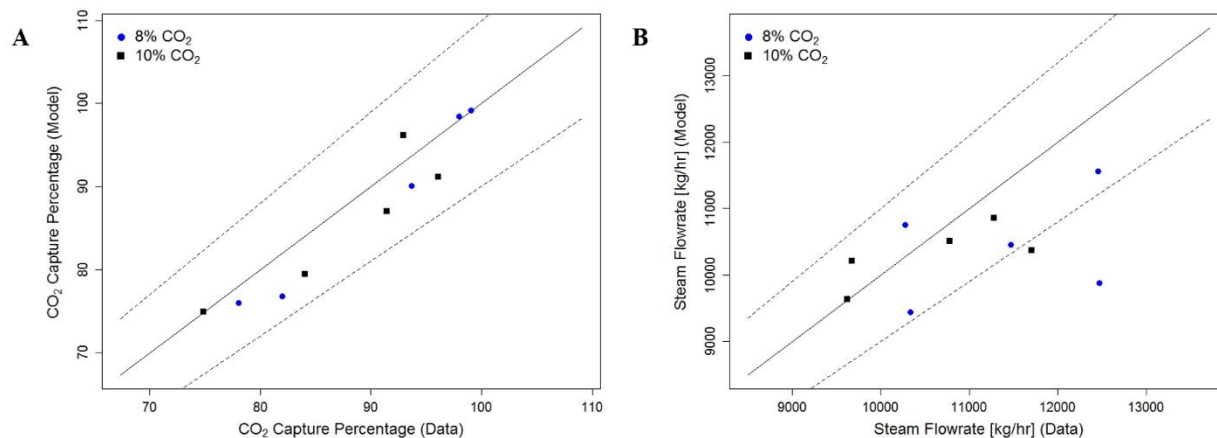


Fig. 4. Parity plots for (A) CO₂ capture percentage and (B) steam flowrate required for test runs performed in Phase 2. Dashed lines represent $\pm 10\%$ error.

As shown in Figure 4, the accuracy of the model for the data collected in Phase 2 is comparable to that in Phase 1. The average percent error for the CO₂ capture prediction and steam requirement prediction is $-2.40 \pm 3.27\%$ and $-5.28 \pm 8.00\%$, respectively. The percentage error for the steam requirement prediction is substantially lower for the data collected in Phase 2 than in Phase 1 due to the absence of test runs with very low ($< 85,000$ kg/hr) solvent flowrate. Therefore, the model was shown to be sufficiently accurate for the region of the input space likely to be economically optimal.

Phase 3 of the test campaign was focused on data collection in regions where the stochastic model predicts relatively high uncertainty for the absorber CO₂ capture percentage. These data, along with those collected in Phase 2, were used to update the mass transfer and interfacial area model parameter distributions. The test matrix for Phase 3 is shown in Table 4.

Table 4. Test matrix for Phase 3 design of MEA test campaign at TCM (First Iteration)

Test	Rich Solvent Flowrate (kg/hr)	Flue Gas Flowrate (Sm ³ /hr)	Steam Flowrate (kg/hr)	CO ₂ in Flue Gas (vol%)	CO ₂ Capture Percent Estimate
3A	133,900	62,600	11,600	8	85.9
3B	115,400	62,300	10,700	8	81.3
3C	111,900	59,100	11,100	8	89.3
3D	120,200	56,100	10,100	8	84.1
3E	119,500	55,000	9,900	8	83.6
3F	81,500	51,100	10,300	8	90.1
3G	57,500	42,500	8,700	8	81.8
3H	39,300	30,800	6,600	8	80.0
3I	48,300	30,400	8,200	10	80.0
3J	85,600	33,800	7,500	10	85.5
3K	103,100	43,000	9,200	10	82.2

The data collected in Table 4, along with case 2A from Table 3 were used in the Bayesian inference procedure based on Eq. 2. In this work, the parameters contained in θ included the leading coefficients for the interfacial area and mass transfer submodels developed in previous work [3], and the parameters contained in θ^* included the thermodynamic model parameters for which distributions were estimated in previous work [2]. Upon obtaining the updated parameter distributions, the refined stochastic model was used to develop a new test matrix, shown in Table 5.

Table 5. Test matrix for Phase 3 design of MEA test campaign at TCM (Second Iteration)

Test	Rich Solvent Flowrate (kg/hr)	Flue Gas Flowrate (Sm ³ /hr)	Steam Flowrate (kg/hr)	CO ₂ in Flue Gas (vol%)	CO ₂ Capture Percent Estimate
3L	96,100	41,300	9,900	10	89.4
3M	94,000	43,500	11,000	10	88.8
3N	119,500	46,500	10,900	10	86.4
3O	150,300	48,200	11,500	10	85.2
3P	130,200	58,400	10,500	8	81.9
3Q	99,900	53,400	10,500	8	90.8
3R	80,800	51,600	12,900	8	88.3
3S	127,500	50,800	10,000	8	88.7
3T	121,200	49,300	9,200	8	85.2
3U	98,200	47,800	8,400	8	81.6
3V	125,500	47,000	9,900	8	94.2

The data collected from the test plan given in Table 5, along with cases 2B-2I in Table 3, were used in a second iteration of the SDoE procedure to update the parameter distributions again. Parity plots for the model prediction of CO₂ capture percentage in the absorber and steam requirement in the stripper for all data collected in both iterations of Phase 3 are given in Figure 5.

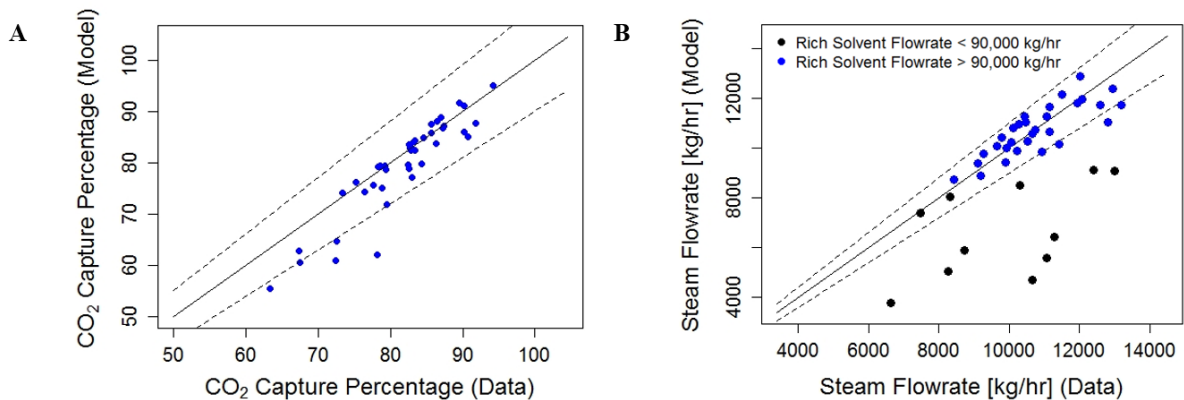


Figure 5. Parity plots for (A) CO₂ capture percentage and (B) steam flowrate required for test runs performed in Phase 3. Dashed lines represent $\pm 10\%$ error.

The average percentage error values for the model predictions of the data collected in Phase 3 are $-2.91 \pm 5.27\%$ for CO₂ capture percentage and $-8.53 \pm 17.20\%$ for the steam flowrate. As with the data collected in Phase 1, there is greater discrepancy in the stripper model for cases in which solvent flowrate is low; the average percentage error in the steam requirement is $-31.05 \pm 17.81\%$ for cases in which the solvent flowrate is below 90,000 kg/hr and $-0.27 \pm 6.04\%$ when it exceeds 90,000 kg/hr. As previously suggested, the underprediction in steam flowrate is likely due to operation inefficiency of the RFCC stripper caused by solvent maldistribution, as the process is operated at much lower solvent flowrate than the stripper design condition.

The probability density functions of the mass transfer and interfacial area parameters, including the prior and posterior distributions obtained after each SDoE iteration, are given in Fig. 6.

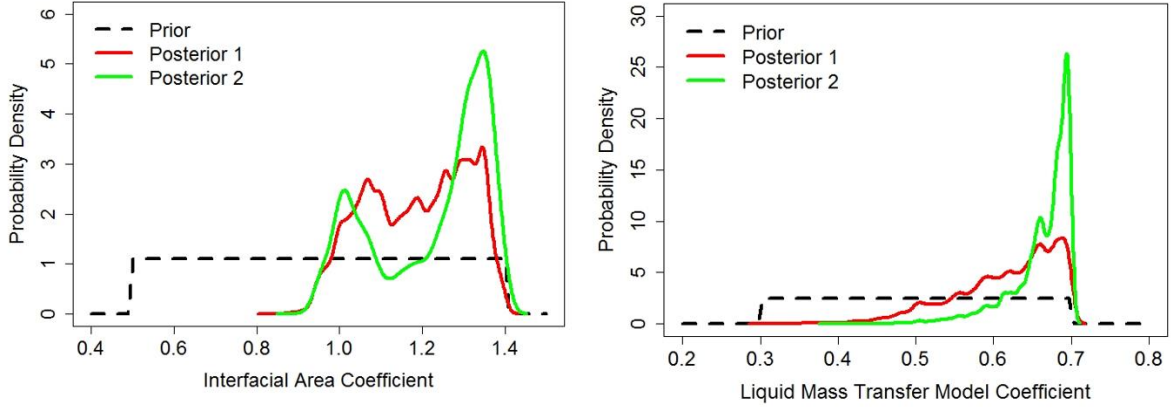


Figure 6. Comparison of prior and posterior distributions of interfacial area and mass transfer model parameters

Uniform prior distributions were initially chosen for the two parameters. The parameter space of plausible values was significantly reduced after incorporating the experimental data from the first iteration of SDoE into the stochastic model through Bayesian inference, with less reduction in the second round of SDoE. The effect of SDoE on model uncertainty reduction is more apparent when considering the model output, namely the CO₂ capture percent in the absorber. The effect of the first iteration of SDoE on reducing model prediction of uncertainty in CO₂ capture percentage is shown in Fig. 7.

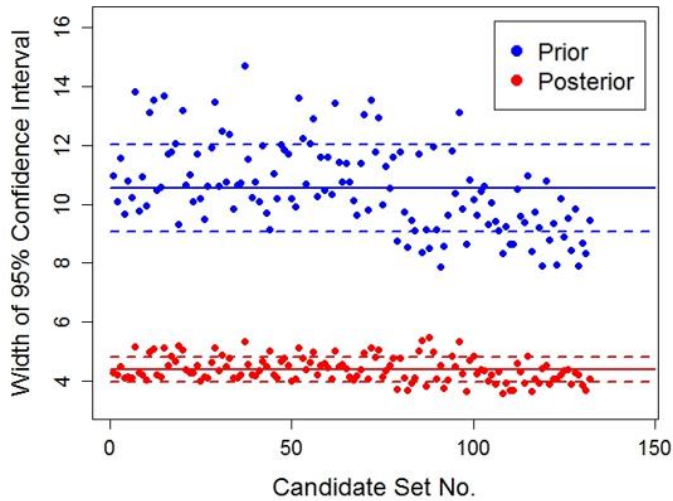


Figure 7. Effect of first round of Bayesian inference on CO₂ capture prediction confidence interval for individual points in candidate set

For the stochastic model prediction using the uniform prior distributions, the average confidence interval width for the CO₂ capture percentage was approximately 10.5% (denoted in Figure 7 by solid line) with standard deviation 1.5% (denoted by dashed lines). For the stochastic model prediction with the posterior distribution obtained after the first iteration of SDoE, the average confidence interval width was approximately 4.4% with standard deviation 0.4%. No further significant reduction in the predicted uncertainty in CO₂ capture percentage was demonstrated in the second round of SDoE. In Figure 7, the candidate set number refers to an index representing a unique combination of input variables (liquid and gas flowrates, CO₂ loading, and CO₂ fraction in flue gas). The percentage of reduction in uncertainty for a given point ($\mathbf{x}^{(t)}$) in the candidate set is calculated as:

$$\text{Percent Reduction} = 100\% \times \frac{[CI]_{\mathbf{x}^{(t)}}^{\text{initial}} - [CI]_{\mathbf{x}^{(t)}}^{\text{final}}}{[CI]_{\mathbf{x}^{(t)}}^{\text{initial}}} \quad (6)$$

where $[CI]_{x^{(i)}}^{initial}$ and $[CI]_{x^{(i)}}^{final}$ represent the 95% confidence intervals in the model prediction of CO₂ capture percentage before and after updating the parameter distributions through Bayesian inference, respectively. For the entire candidate set, the average percent reduction in the uncertainty is $58.0 \pm 4.7\%$, which is comparable to the reduction in the previous SDoE-based test campaign executed at NCCC [5,6]. As the ability of the SDoE methodology to reduce parametric uncertainty in a process model for an aqueous MEA system has been demonstrated in multiple campaigns, it may be considered a promising technique for designing future test campaigns to effectively increase fundamental understanding of novel CO₂ capture systems.

4. Conclusions and Future Work

In summary, a sequential design of experiments methodology was implemented for executing a test campaign for aqueous MEA at TCM, guiding collection of process data to refine the parameter distributions in the stochastic process model. This resulted in an average reduction of around 58% in the uncertainty in the prediction of CO₂ capture percentage. The deterministic model, or the model without parameter uncertainty, also predicted the plant performance accurately, with an average error in percentage of CO₂ capture of $-2.74 \pm 4.47\%$ for the first three phases and an average error of $-8.52 \pm 14.85\%$ for the reboiler steam requirement. An exception to the accurate performance of the model is for data collected under impractical operating conditions (low solvent circulation rate, in which solvent maldistribution in the stripper column was noted). For data collected when the system was operated with rich solvent flowrate below 90,000 kg/hr, the percent error in the reboiler steam prediction was $-23.92 \pm 15.70\%$. However, the percentage error in the steam prediction is $-1.17 \pm 6.65\%$ for data collected when the rich solvent flowrate was above 90,000 kg/hr. The insights gained during the execution of SDoE guided the development of a new SDoE module with capability for straightforward implementation of the aims used in this experiment [16] that has been implemented in the Framework for Optimization, Quantification of Uncertainty, and Surrogates (FOQUS). This is available as part of the aforementioned CCSI Toolset and will enable the SDoE process to be implemented in a more streamlined manner in future applications. In planned future work, the CCSI² team will apply the SDoE methodology to novel CO₂ capture technologies with the primary goal of refining initial process models by reducing their uncertainty, and thus the inherent risk associated with preliminary models of new processes, through guided data collection.

Disclaimer

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